



Congressional Agenda Control and
the Decline of Bipartisan Cooperation

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Abstract

Evidence shows that polarization in Congress has increased substantially since the 1970s while polarization in the public has increased much less, if at all. These two patterns of polarization suggest that responsiveness by members to their constituents has declined. By breaking apart congressional behavior, however, this paper suggests that this presumption is misleading. Looking at House roll call votes and bill cosponsorship coalitions, this paper suggests that although partisan behavior has increased substantially in roll call votes, the same is not true for bill cosponsorship coalitions. In turn, this suggests that while responsiveness has declined when considering roll call voting, responsiveness has increased when considering cosponsorship coalitions. These divergent patterns can be reconciled by taking into consideration congressional agenda control. As congressional parties sorted, partisan legislation became increasingly likely to face a roll call vote whereas bipartisan legislation became less likely to face a roll call vote.

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There is a presumption in the literature – both by journalists and academics – that party polarization has contributed to declining responsiveness by members of Congress to the public. Evidence shows that polarization in Congress has increased substantially since the 1970s while polarization in the public has increased much less, if at all. From here, the assumption is made that there is a disconnect between voters and their representatives (Fiorina and Levendusky 2006), and that responsiveness has declined. This finding is troubling from both a normative perspective and a rational choice perspective, where the former suggests flaws in representation and the latter suggests that elections are failing to produce members whose preferences match that of the median voter. By breaking apart congressional behavior, however, this paper indicates that this presumption is misleading. Looking at House roll call votes as well as bill cosponsorship coalitions, this paper finds that although partisan behavior has increased substantially in roll call votes, the same is not true for bill cosponsorship coalitions. In turn, this suggests that while responsiveness has declined when considering roll call voting, responsiveness has increased when considering cosponsorship coalitions. These divergent patterns can be reconciled by taking into consideration congressional agenda control and the selection of bills to receive roll call votes. By bringing an explicit focus to the congressional roll call agenda, this paper examines how party control over the congressional agenda, and specifically the temporal changes in the selection of bills to receive roll call votes, affects the degree of bipartisan cooperation in the House of Representatives.

These results have a number of implications for understanding polarization in American politics, agenda control in congressional policy-making, and the importance of carefully considering the potential biases in roll call data. An exclusive focus on roll call votes to measure bipartisanship or party polarization may be misleading and can either under or over-estimate

partisanship, depending on how the congressional agenda is structured. While it is clear that voting patterns have changed since the 1970s, consideration of the agenda and the selection of bills to face roll call votes must be taken into account, rather than taking voting patterns as exogenous.

Literature on political parties has often debated when and how parties are influential, particularly in weak party systems such as the United States (Cox and McCubbins 2005; Krehbiel 1993). The findings in this paper are similar to the findings by Cox and McCubbins, suggesting that party influence occurs through congressional agenda control. Rather than focusing on the ultimate degree of party unity on votes, however, this paper focuses on the changing likelihood over time that bipartisan bills receive roll call votes. From this perspective, changes in the construction of the agenda over time show an increasingly partisan strategy. Similar to Fiorina's (2004) finding of similarities in public opinion and values between citizens in red and blue states, my findings suggest that this common ground may extend to members of Congress to a greater degree than is generally recognized. When we assess legislative behavior absent agenda control, members of Congress continue to engage in bipartisanship at rates similar to those thirty years ago. Members of Congress have become increasingly representative of their districts when using bill cosponsorship coalitions, despite the fact that representation has declined when using roll call votes.

This paper proceeds as follows. The next section sets out the groundwork for the standard argument that party polarization has increased, bipartisan cooperation has declined, and thus, that electoral responsiveness has declined. The third section breaks apart legislative behavior from the typical reliance on voting patterns, considering members' bill cosponsorship coalitions. The fourth section reconciles the divergent patterns of bipartisan cooperation in voting and bill

cosponsorship by considering the strategic use of agenda control. The final section concludes.

Party Polarization and the Decline of Electoral Responsiveness

Half a century ago, political scientists yearned for distinctive, responsible parties (APSA 1950) but the resurgence of polarized parties since the 1970s has left few satisfied. Though there are some that continue to defend partisanship and polarization (e.g., Muirhead 2006), the more common refrain is that party polarization has come at the expense of an idealized era of bipartisanship (e.g. Eilperin 2006). Combined with evidence that the public has polarized little, if at all (Fiorina et al. 2004; DiMaggio et al. 1996), this body of research suggests that there is a disconnect between the mass public and the government (Fiorina and Levendusky 2006). Ultimately, this suggests that collective or aggregate electoral responsiveness has declined. This section summarizes the prevailing argument, drawing on the voting patterns of members.¹

Increasing Polarization and Declining Bipartisanship

The standard measures of polarization as well as legislative behavior come from the analysis of roll call votes in Congress. This section considers various measures of partisanship and bipartisan cooperation, all of which rely on roll call votes themselves or on the subsequent ideology estimates derived from votes. Regardless of the ideology estimate that is used (NOMINATE, Americans for Democratic Action (ADA), etc.), party polarization – measured as the difference in party means or medians - has increased significantly, and almost monotonically, since the 1970s.

Turning to the prevalence of partisanship in Congress, let us look at the incidence of party unity relative to bipartisan votes in Congress. Congressional Quarterly (CQ) defines party unity votes as roll call “votes that split the parties, a majority of voting Democrats opposing a

¹ The analysis focuses on the period from 1973 to 2004. This choice was driven by the adoption of modern roll call voting procedures, and the use of cosponsorship in the House being restricted prior to the late-1960s.

majority of voting Republicans” (“CQ Fact Sheet Bipartisan Voting” 1970, 1139). Since the discussion of cosponsorship coalitions will focus on the frequency of bipartisan behavior, it is useful to consider the flip side of party unity votes, or what CQ classifies as bipartisan votes. Bipartisan votes are “roll-call votes on which a majority of voting Democrats and a majority of voting Republicans agreed” (“CQ Fact Sheet Bipartisan Voting” 1970). The CQ Bipartisan Measure (see Figure 1) for the House of Representatives shows a fairly steady decline in bipartisanship from the 1973 to 1995 with some recovery of bipartisanship beginning in the 105th Congress (post-1996), but declining again after 2001.²

Since the CQ measure of bipartisanship looks at what majorities of the two parties are doing, it may miss additional and important variation. In particular, the CQ measure misses the extent of skew within voting (i.e., is it 90% of one party against 90% of the other party or 52% of one party against 52% of the other). To parse out this element in the analysis of bipartisanship, I calculate a ‘Difference in Party Support Score’ for each vote, focusing on the Yea votes.³ For each vote, I calculate the percent of voting Democrats voting Yea and subtract the percent of voting Republicans voting Yea. The resulting score ranges from -100 to 100, where -100 indicates a vote on which no Democrats voted Yea and all of the Republicans voted Yea. The midpoint of 0 indicates bills where the same percentage of Democrats and Republicans voted Yea. Thus, the ends of scale reflect the greatest partisanship and the middle represents the

² A similar pattern is found when looking just at final passage votes and omitting procedural issues. A number of scholars have noted that procedural votes have become increasingly partisan. To check whether the relationship between polarization and bipartisanship holds when procedural votes are excluded, I draw on data compiled by David Rohde (2004) for House roll call voting. Looking only at final passage votes and only those on House bills (rather than resolutions or amendments), I find that the percentage of roll call votes that are bipartisan (using the CQ definition) declined from approximately 80 percent in 1973 to 40 percent by 1994. Although there was an increase in bipartisanship under the Republicans, the overall trend of the series has been a strong decline in bipartisanship consistent with the aggregate roll call results. Plots of the trends using final passage votes are available in the online appendix.

³ The analysis focuses on Yea votes since they are most similar to cosponsorship coalitions in indicating support for a piece of legislation.

greatest bipartisanship, assessed as where voting divisions are the same within each party.

I first plot the density of the scale by Congress for Yea votes when all House roll calls are included (Figure 2).⁴ The trend over time has been an increase in density at the two poles of the distribution at the expense of the middle, or of the most bipartisan bills.⁵ In sum, bipartisan cooperation on roll call voting shows significant declines across the last three decades, particularly from the 1970s through the mid-1990s. This decline in bipartisanship is consistent with the conventional story of the rise in elite polarization.

Polarization in the American Public?

Despite the abundance of evidence that Congress has become more polarized and partisan, there is significantly more skepticism that the public has grown as polarized. For instance, DiMaggio et al. (1996) found little evidence that the American public's social attitudes had become more polarized over the previous two decades, with the exception being attitudes on abortion. Similarly, Fiorina (2004) found little evidence of growing disparities between the opinions of red and blue state voters. While political elites – including party activists – have polarized, the general electorate has not shifted its positions. As noted by Levendusky (2009), the public has sorted but has not polarized. The electorate only appears polarized because it is choosing between the alternatives set by the (polarized) political elite (Fiorina et al. 2004). While some scholars have contended that the public has polarized (e.g. Abramowitz and Saunders 2008), these findings focus more on the choices presented to the public, than to polarization in their underlying preferences (Fiorina et al. 2008).

⁴ This plot omits unanimous and near unanimous (90% in favor) votes, as is done in NOMINATE calculations as well. When unanimous and near unanimous votes are included, the majority of votes are bipartisan. This further suggests that roll call vote analyses, which tend to omit unanimous and near unanimous votes because they do not provide additional information for distinguishing between the preferences of various members, are biased against finding bipartisan cooperation.

⁵ Restricting the analysis to final passage votes on House bills also shows a decline in bipartisanship, though patterns are skewed toward the majority party (see online appendix for comparable figure).

Polarization and Declining Responsiveness

Although political scientists often argue that concerns about reelection drive members to be representative of and responsive to their constituents, the dual patterns of increased elite polarization and little mass public polarization suggest that members are achieving reelection without being representative of their constituents. This breakdown in representation can be seen both in the relationship between legislators' voting scores and summary measures of district preferences, as well as from survey responses regarding cooperation and bipartisanship.

Scholars have long argued that reelection is a driving, if not *the* driving, force behind the behavior of members of Congress (Mayhew 1974). Normatively, the 'electoral connection' suggests that both sides win; members of Congress get what they want (to stay in office), and the public gets what they want (responsive representatives). Taken to the extreme, Anthony Downs (1957) argued that in a two-party system, both candidates should converge to the median voter in the district.

A large body of work has grown around the question of representation and whether members converge to the position of the district median – what would be perfect representation (e.g., Ansolabehere et al. 2001; Burden 2004). Despite findings that show that candidates rarely fully converge in their positions, scholars have maintained that the median voter is important and that members can be electorally punished if they stray too far from the districts' preferences (e.g. Canes-Wrone et al. 2002; Carson et al. 2010).

Additional evidence that the public wants members who are less polarized, and who are more bipartisan, comes from survey work. A May, 2010 poll by Pew found that “a substantial minority (42%) say they would be more likely to vote from a candidate who will make compromises with people they disagree with; only about half as many (22%) say they would be

less likely to back a candidate willing to compromise” (Pew 2010). Similarly, Carson et al. (2010) find that potential candidates lose support when respondents are told the candidate is very partisan.

Logically, it appears that an increasingly partisan Congress without equivalent polarization in the public damages representation. What does the data suggest? To begin, let’s select a measure of district preferences, with which to judge legislative behavior in relation to. This paper draws on the *Normal Presidential Vote* in the district, measured as the mean two party presidential vote in the previous two elections by the party of the Representative (Canes-Wrone et al. 2002).⁶ For instance, if the member is a Republican I use the mean Republican presidential vote in the last two presidential elections and if the member is a Democrat I use the mean Democratic presidential vote in the last two presidential elections.⁷ The presidential vote has been found to be an excellent proxy of district-level partisanship (Levendusky et al. 2008).

There are a number of ways to look at legislative behavior for evidence of electoral responsiveness. In this section I draw on party unity support scores. The *Party Unity Support Score* calculates the percent of times that a member votes with his or her party on party unity votes (defined when a majority of Democrats vote against a majority of Republicans). As mentioned in previous sections, CQ classifies votes as either party unity votes or bipartisan

⁶ I use the normal presidential vote rather than the member’s own past vote share because of concerns with foresight on the part of the legislator as well as incumbency advantage. If a member foresees the importance of appearing as bipartisan or partisan, his vote share is likely to reflect that. In effect, assuming any foresight by members means that their vote share, even at time $t-1$, is endogenous to their behavior. An additional problem is the personal vote for incumbents. That is, a large margin of victory may imply a partisan district or it may imply that a member has created a strong personal brand, potentially because of their responsiveness to the district.

⁷ In all analyses in this paper, all districts (where data is available) are included. This includes districts that were redistricted. Ideally, instances where a district boundary was redrawn would be omitted, and the normal presidential vote in subsequent years would be omitted to drop the presidential year in the first year after redistricting. However, the author’s current data on redistricting only includes an indicator for whether a district is in a state that was redistricted, meaning that removing these cases leaves minimal observations in the 93rd (1972), 98th (1982), 103rd (1992), and 108th (2002) Congresses. As a result, redistricted cases are included and are a source of measurement error.

votes. Thus, the logic of this measure, applied to the argument that members are electorally responsive, is that members in competitive districts should have lower party unity support scores than members in safe districts. The drawback of this score is that it does not capture the relative frequency of party unity votes in general.

Looking at the correlation between the normal presidential vote in the district and the party unity support score yields suggestive evidence of changes in representativeness over time (see Table 1). Although initially showing a relatively strong correlation with the normal vote, party unity support scores show a weakening relationship from the 97th to 102nd Congresses.

A second approach further parses out these patterns. I use a quasi-binomial model of legislative behavior, which accounts for the likelihood that a member engages in partisan (or bipartisan) behavior relative to the number of times they do not. ‘Successes’ in the model are votes with the party. The primary independent variable is the *Normal Presidential Vote* in the district. I allow this effect, as well as the intercept, to vary by Congress. Finally, I include a number of individual level covariates – member of the *Majority Party*, *Female*, *Age*, *Tenure*, and an indicator for whether the member holds a *Leadership* post (Speaker, Majority or Minority Leader, or Whip). For each of these control variables, the effects are constrained to be constant across time. A quasi-binomial, rather than a binomial model, is used in order to allow for over-dispersion.

To assess the degree of responsiveness to the normal presidential vote over time I examine the predicted probability of partisan support over the range of the normal presidential vote for each Congress. The calculation of the predicted probability holds all variables in the model except the *Normal Presidential Vote* at either their mean or median. Figure 3 presents the predicted probability of a member voting with his party on party unity votes for selected

Congresses.⁸ The 95% confidence intervals are included in the dashed lines. Over time, we see both an intercept shift toward greater party support as well as a decrease in the slope, indicating less differentiation across members. By this measure, the low point for representation occurs in the 101st Congress.

It is important to note that neither of these approaches provides a definitive benchmark of ideal representation. Rather, both of these approaches rely on comparisons over time within the measure, looking at either changes in the correlation or changes in the intercept and slope of the estimated relationship. In making the argument that these changes reflect declining representation, I am making the assumption that a higher correlation between the measures or a lower intercept and steeper slope reflect higher levels of representation. With these caveats in mind, both of these roll call patterns are suggestive that electoral responsive has declined.

Breaking Apart Legislative Behavior - Bill Cosponsorship Coalitions

Although roll call votes are the standard measure of legislative behavior, and often for good reason, it is important to remember that roll call votes are not a random sample of legislation. Therefore, this paper considers a second measure of legislative behavior – bill cosponsorship coalitions – to understand patterns of partisan and bipartisan cooperation in Congress. Of course these are not the only two forms of bipartisan cooperation that can occur in Congress. Bipartisanship may also be observed in the processes used in the House (i.e., the use of closed versus open rules or the composition of conference committees), or even in the rhetoric of members' speeches. Nonetheless, roll call votes and bill cosponsorship coalitions are two areas that provide a unique opportunity to both systematically examine decisions by all members of Congress and to examine the influence of congressional agenda control.

⁸ The pattern over time is consistent in the Congresses not presented here and graphs of all Congresses are available upon request. The full model specification is available in the online appendix.

Despite the prevalence of roll call votes in Congressional literature, the bills that reach roll call votes are not a random sample of the bills or issues that have been brought up in Congress. First, many scholars have argued that party leaders strategically put up issues for roll call votes that divide the parties from one another but do not show divisions within their own party (Poole 2004; Carrubba et al. 2008). If roll call votes are more likely on legislation that pit the parties against one another, this has the potential to bias analyses of roll call votes from finding high levels of bipartisanship. This is particularly true if legislation with bipartisan support passes via voice votes rather than roll call votes. Second, the likelihood that legislation, and particularly bipartisan legislation, faces a roll call vote may not be constant across time. For instance, Lynch and Madonna (2008), find that the prevalence of voice votes relative to roll call votes on significant legislation varies both across time and across issues. If the selection of bills for roll call votes is biased against bipartisan legislation, or if the selection of bills for votes varies over time in a way that is related to whether legislation has bipartisan support, then both the level of bipartisanship and the decline in bipartisanship over time may be misleading.⁹

Because of these concerns, I explore bill cosponsorship coalitions as a second measure of bipartisanship in Congress. As noted by Kessler and Krehbiel (1996), cosponsorship is less likely to be subject to agenda control. Like roll call votes, however, bill cosponsorship coalitions allow all members the opportunity to take a position and side with a coalition of other members. By looking at bill cosponsorship coalitions, I am able to assess bipartisanship absent the strategic considerations of which bills face roll call votes.

Bipartisanship in Cosponsorship Coalitions

Bill sponsorship has long been understood as an effective and relatively easy way for

⁹ Unfortunately, little research has systematically examined voice votes so this work, like others, focuses on whether legislation faces a roll call vote and defines the agenda as such.

members of Congress to become involved in the policy process. Similarly, cosponsoring legislation carries a number of possible benefits, both within Congress and among constituents. As such, cosponsorship has become a frequent activity in both the House and Senate in recent decades.

Since it was first allowed in the House in 1967, cosponsorship has largely replaced the use of duplicate bills (Thomas and Grofman 1993), suggesting that members view cosponsorship in much the same way as they do sponsorship because they can make the same claims back in their constituencies. There are a number of reasons to believe that cosponsorship coalitions are a useful stage in the legislative process to assess bipartisanship. First, cosponsorship signals the political benefits of a bill (Koger 2003, 227). Second, what a member cosponsors, and with whom they cosponsor, allows members to send messages to their constituents. “Cosponsoring helps clarify your message. That way people know where you are... and that trickles down to constituents” (quoted in Koger 2003, 232).

Though there have been debates about why members cosponsor, the assumption in this paper is that members cosponsor legislation that they genuinely favor (similar to a preference or matching argument). However, since members only cosponsor a fraction of the entire number of bills that they might favor, the selection of which bills to cosponsor (among those bills that they like) is assumed to reflect concerns for signaling (either to constituents or internally to Congress) or reciprocity among members.

To measure bipartisan cosponsorship, I utilize cosponsor data collected by James Fowler (2006) that provides a matrix of all bills and cosponsors in a given Congress. I use this raw data

to create bill level measures of bipartisanship for all House bills.¹⁰ Before focusing on bipartisan cosponsorship, it is important to understand the trends in more general cosponsorship patterns. On the whole, the use of cosponsorship has increased in the House since the 93rd Congress (see online appendix for summary statistics). Whereas only one-quarter of House bills were cosponsored in the 93rd Congress, since the 99th Congress, more than half of all bills have been cosponsored. The median number of cosponsors on a House bill has also risen over time. Among bills that are cosponsored, the median number of cosponsors has nearly doubled in the period of analysis, from six in the 93rd Congress to eleven in the 105th through 108th Congresses. However, members are selective in their cosponsorship. No more than six percent of all cosponsored bills have more than a hundred cosponsors.

Since cosponsorship has not been analyzed as extensively as roll call votes, there are no standard measures of partisan or bipartisan cosponsorship. Therefore, this paper considers a number of possible specifications. I begin by creating a dichotomous measure of bipartisanship where bills are bipartisan if they are above the lower bound of bipartisanship on the scale and are partisan otherwise. Taking the simplest definition of bipartisanship – all cosponsored bills that are not uni-partisan (where uni-partisan refers to bills where all of the cosponsors are from the same party as the sponsor) – suggests that bills have been and continue to be overwhelmingly bipartisan, as over eighty percent of bills meet this criteria of bipartisan.

To create a more stringent definition, I restrict my definition of bipartisanship to those bills that have at least twenty percent of the cosponsors from the party other than the original

¹⁰ I use only House bills throughout the analysis, omitting resolutions and amendments. Although all House bills are included in the analysis, omitting commemorative legislation, which coding is available for through 2002, does not change the results. The measures of bipartisanship with and without commemorative legislation are correlated at 0.98.

sponsor (see Figure 4).¹¹ When all House bills are considered, the proportion of bills that are bipartisan increases over time, largely due to the overall increase in cosponsorship. When the analysis is restricted to only cosponsored bills, I find that although there has been some movement and a slightly downward trend in bipartisanship over time, the magnitude of the change is surprisingly small. Over the entire period of the 93rd to 108th Congresses, the range of bipartisan cosponsorship is between 47% and 61%. The low point of 47% occurs in 1975, 1993, and 2004, which does not fit with patterns of bipartisanship by standard roll call accounts. Though the interpretations of the exact percent of bipartisan votes and bipartisan cosponsorship coalitions are not comparable, given the differences in definitions and cut-points, the key insight is the disparate patterns over time within the two measures.¹²

As done in the analysis of voting, I also create a difference in party support score for each bill. Since cosponsorship occurs only on the positive side, this is similar to looking at the Yea votes. For each cosponsored bill, I subtract the percent of Republican cosponsors from the

¹¹ Additional definitions of bipartisanship, including a raw number of cosponsors from the opposing party (5, 10, and 20) as well as different percentages of cosponsors from the opposing party (20%, 30%, and 40%) have also been used for comparison. In all cases the pattern over time is similar with the primary difference being a shift in the intercept. A potential concern about the analysis presented above is that those bills classified as bipartisan may not be that different from those bills that are classified as partisan, particularly since the inferences are often made from a small number of cosponsors. To verify that those bills classified as bipartisan are distinct from those bills classified as partisan in the analysis, I focus just on those bills that receive a roll call vote and examine the probability that each type of bill – partisan and bipartisan by cosponsorship - receives a bipartisan roll call vote. That is, of all bipartisan (or partisan) cosponsored bills that face a roll call vote, what proportion end up having a bipartisan roll call vote (as defined by the CQ measure)? Using data from the Policy Agendas Project, Rohde's dataset of House roll call votes, and my bipartisan bill cosponsorship measures indicates that between one-quarter and one-half of bills with bipartisan cosponsors that reach roll call votes result in a bipartisan vote. The average for 1973 through 2000 (the last year in which all three data sources are available) is just about one-third. In contrast, between one-twentieth and one-fifth (with an average of three-twentieths) of bills with partisan cosponsors that reach roll call votes result in a bipartisan vote. Although it is not impossible for partisan cosponsored bills to result in a bipartisan roll call vote, it is rare. In all years, bills with bipartisan cosponsorship coalitions are more likely to result in a bipartisan roll call vote than bills with partisan cosponsorship coalitions. This suggests that the cosponsorship measures are capturing important variation, and that the importance of this variation extends to voting patterns of the chamber as a whole.

¹² Using a continuous measure of the percent of cosponsors from the party opposite the party of the sponsor, reiterates the persistence of bipartisanship across time. This measure is similar to the party support score except that it directly accounts for the party of the bill's sponsor. Across the period of analysis, the mean percentage of bipartisan cosponsors fluctuates around 30%, indicating that, on average, one third of a bill's cosponsors are from the party opposite the party of the original bill sponsor.

percent of Democratic cosponsors. Like the roll call-based measure, the resulting the scale ranges from -100 to 100, where the endpoints reflect bills with all Republican cosponsors and all Democratic cosponsors, respectively.

Looking at the density of the difference in party support scale for each House of Representatives from the 93rd to the 108th Congress indicates that bipartisanship persists. As seen in Figure 5, although the distribution of cosponsor coalitions is bimodal in the sense that there are a large number of both Democratic and Republican partisan bills, the middle (i.e., the most bipartisan bills) has not vanished over time. The primary change in the distribution of bills is that it has become less lopsided toward the Democratic side, particularly once the Republicans gained majority status. Whereas partisan bills used to be predominantly Democratic, there is a relative parity of Democratic and Republican partisan bills in more recent Congresses and a growth in moderately bipartisan bills on the Republican side. More on point for this analysis, however, is the finding that the center of distribution persists across time. If there is a decline in bipartisanship, we should have seen the density of both tails grow at the expense of the center, creating a greater U-shaped pattern over time.

Regressing these various measures of bipartisanship on time reiterates the decline of bipartisanship on roll call votes but a greater persistence of bipartisanship on bill cosponsorship.¹³ Beginning with the CQ measure of all roll call votes, the estimated coefficient on time is -0.45 ($p < 0.05$). The magnitude of the effect increases when only final passage votes are included, to -0.85 ($p < 0.001$). In contrast, for the 20% definition of bipartisan bill cosponsorship, the estimated coefficient on time is -0.25 ($p < 0.01$). Similar patterns are found when bipartisanship is measured using the difference in party support scales. In this case, we

¹³ Full model specifications available in the online appendix. Time is measured as the number of years since 1973.

expect the direction of the coefficient to be positive since lower absolute values of the scale reflect greater bipartisanship. Defining bipartisanship as the mean of the absolute value of the party scale, we get an estimated coefficient on time of 1.16 ($p < 0.001$) when looking at yea roll call votes (coefficient of 1.11, $p < 0.001$ for final passage votes), but a coefficient of -0.02 ($p = 0.72$) when looking at bill cosponsorship coalitions. Thus, regardless of the technique for estimating bipartisanship, bipartisan cooperation shows a greater decline over time in roll call votes than in bill cosponsorship coalitions.

Cosponsorship Coalitions and Electoral Responsiveness

At the outset of this paper, I argued that standard accounts of party polarization often imply a decline in electoral responsiveness. In a previous section, I empirically demonstrated this decline, drawing on the party unity support scores of members on roll call votes. In this section, I return to the assessment of responsiveness, following the same approaches but using bill cosponsorship coalitions rather than party unity support scores on votes. As in the previous section, district preferences are measured using the *Normal Presidential Vote* in the district. Focusing on bill cosponsorship patterns for the dependent variable, I define a member's propensity for *Bipartisan Cosponsorship* as the percentage of bills they cosponsored that are bipartisan. Bipartisan is defined using the definition where at least twenty percent of the cosponsors are from the party opposite the party of the sponsor.

Looking at the correlation between the normal presidential vote in the district and bill cosponsorship behavior offers a stark difference from the patterns of roll call voting. As seen in Table 1, the normal presidential vote and members' party unity votes showed a weakening relationship over time. In contrast, bipartisan cosponsorship measures show little to no relationship for the first part of the series, but a growing correlation since the 102nd Congress.

Since the 103rd Congress, the correlation between the normal vote and members' bill cosponsorship coalitions has generally been larger in magnitude than the correlation between the normal vote and members' party unity support scores.

A second approach further parses out these patterns, again using a quasi-binomial model. In this model, 'successes' are bipartisan cosponsorship coalitions. The primary independent variable and control variables are the same as in the analysis of roll call votes. Figure 6 presents the predicted probability of cosponsoring a bipartisan bill for a selected set of Congresses. As expected, the slope of the line is in the opposite direction than in the previous table since we have changed to looking at bipartisanship (rather than partisanship). More important, however, is the different pattern that we observe over time. Here, we see little movement in the intercept over time and an increase in the steepness of the slope. Both of these patterns suggest an increase in responsiveness to district preferences when looking at bill cosponsorship coalitions.

Whereas members' responsiveness to their districts declined in the roll call-based measure, it increased over time when using the bill cosponsorship measure. This result has significant implications both for the empirical study of Congress and representation as well as the normative interests in these questions. It is also worth noting that the increase in responsiveness of the cosponsorship measure corresponds time-wise to the low point of the roll call measure. Members of Congress may have realized that their roll call behavior placed them as out of step with their districts (potentially because of the types of bills that were selected to face roll call votes) and thus turned to other forms of legislative behavior to show responsiveness.

Though perhaps thought to be less visible than votes, members' patterns of bill cosponsorship are transmitted to constituents through member websites, newspaper endorsements during elections, among other means. Members discuss their patterns of bill

cosponsorship when filling out questionnaires for newspaper endorsements, often highlighting their bipartisan nature (e.g. Manzullo 2010; Foster 2010). Newspaper editorials then often discuss the centrist, independent, or bipartisan features of members. For instance, in their endorsement of Democratic incumbent Ron Klein, the editorial team at the *Palm Beach Post* noted that, “Rep. Klein has participated in bipartisan legislation to improve veterans’ health care benefits” (“Endorsement: Klein for U.S. House District 22” 2010).

Agenda Control and the Decline of Bipartisanship

Thus far, the data have provided the following empirical patterns – partisanship has increased while bipartisan cooperation has declined when roll call votes are examined, bipartisanship has changed relatively little on bill cosponsorship coalitions, and electoral responsiveness has declined when measured with roll call votes but has increased when measured with bill cosponsorship coalitions. How can we reconcile these divergent patterns? One explanation is that party leaders have increasingly pursued a partisan agenda. This change in agenda content may reflect decreases in overlap between the parties, combined with the benefits that come from creating distinctive party brands. This section presents an overview of this argument and some empirical evidence that supports this claim.

Cox and McCubbins (2005, 5) suggest that the primary tool of the majority party is picking which bills will be voted on at all, and only secondary do they focus on garnering enough votes to get a victory. The party leadership is wary to put up issues for votes that create divisions within their own party. In contrast, the leadership often has incentives to place issues up for votes that are expected to divide the parties, with their party coming out on the winning side. A Congressional Quarterly staff writer notes that:

Republican leaders followed a model they adopted in 2000, in which they carefully orchestrated what came to a vote and managed intraparty conflict to minimize dissent on

the floor. In both the House and the Senate, there were fewer total roll call votes than in 2000, but the percentage of total votes that were party unity votes were somewhat higher in 2004 than in 2000. That is a reflection of the Republican leaders' choreography (Poole 2004).

Political scientist Sarah Binder suggests that "The rules of the game are easy enough to manipulate by a majority party to foreclose opportunities to vote on alternatives that would attract bipartisanship" (quoted in Poole 2004). Electoral strategies geared at motivating the base to turnout can further heighten the disincentive for bipartisan agenda. Adam Nagourney, a reporter for the *New York Times*, noted that when both sides are concerned with motivating their base, the agenda difference between the two is much more dramatic (Nagourney 2003).

At the same time however, the majority party must be cognizant of the need to produce of record of legislative success. Bipartisan legislation has a better chance at becoming law, and major legislation with bipartisan support is often more enduring over the long-term (e.g. Mayhew 1991). Thus, the desires of the party leadership to push programmatic, partisan legislation will be tempered by internal divisions as well as other factors that would limit the success of partisan legislation, including divided government and small majority seat shares. This section focuses on the degree of internal division.

Overtime, the sorting out of the political parties, resulting in reduced overlap between the ideologies of the members, has combined with agenda control to contribute to an increasingly partisan roll call agenda. This sorting is particularly apparent in the South, where Republicans replaced conservative Democrats. Sorting of the two parties, even without having the Democrats move left or the Republicans move right, creates opportunities for partisan legislation. As overlap declines, there is an increase in the number of status quo points for which an alternative exists that receives the support of the majority party median and the floor median, but does not receive the support from the minority party median. That is, there is an increase in the number of

issues that meet Cox and McCubbins' (2005) criteria for being placed on the agenda (support of the majority of the majority party) but that will result in party unity rather than a bipartisan vote.

In order to explore how the incidence of bipartisanship changes when considering the cosponsorship stage versus the roll call vote stage of the legislative process, Figure 7 looks at the percentage of cosponsored bills that are bipartisan under the 20 percent definition for all House bills, for bills that reach roll call votes, and for bills that reach roll call votes *and* are classified as important. Two definitions of important legislation are used. The first uses Congressional Quarterly Almanac (CQ) and the second is original data that was collected matching up Mayhew's important legislation and George Edward's list of important failures with the appropriate bill numbers.¹⁴ In each case, the graph plots the percent of bills in each category that are bipartisan in their cosponsorship coalitions. The lines are smoothed using a loess procedure for ease of interpretation. Here, higher percentages reflect greater bipartisanship. Comparing all cosponsored bills with bills that reach roll call votes, there is a greater decline in bipartisanship in those bills that face roll call votes than there is in overall bipartisan cosponsorship of legislation. Beginning in the mid-1990s, however, there is a resurgence in bipartisanship in bills that reach roll calls, consistent with earlier plots of bipartisan votes. The same pattern is seen when looking only at important bills (CQ) and even more apparent when considering major passages and failures. Combined, these findings suggest that the decline in bipartisanship evident in roll call voting is largely driven by the choice of which bills receive roll call votes.

A second approach to distinguishing how agenda control may influence the measurement of bipartisanship is to look at the conditional probability of reaching a specific stage in the policy

¹⁴ An updated list of Mayhew's important legislative enactments and George Edward's major legislative failures was graciously provided to the author by Sean Theriault. The bill numbers of the specific piece of legislation, as well as duplicate bills and other similar proposals, was coded using CQ Almanac, the Thomas search engine from the Library of Congress, and relevant newspaper articles on the legislation if necessary.

making process given that a bill is either partisan or bipartisan in its cosponsorship coalition. Using the 20% definition of bipartisan legislation, I analyze the conditional probability of reaching a roll call vote given that a bill is bipartisan (or partisan).¹⁵ Bipartisan bills have a greater probability of reaching a roll call vote than partisan bills. However, I find different patterns over time in the conditional probability of reaching a roll call vote. Whereas the conditional probability of a partisan bill reaching a roll call vote slightly increases over time (from 0.01 to 0.04), the conditional probability of a bipartisan bill reaching a roll call vote generally declines from the early 1970s, when it was 0.07, to the mid-1990s, when it was 0.03. The conditional probability that a bipartisan bill reaches a roll call vote increases only after the Republican takeover in the 104th Congress. This suggests that legislation with bipartisan support early in the legislative process (as evidenced by its cosponsorship coalition) was increasingly prevented from reaching a roll call vote from the early 1970s through the mid-1990s.

As mentioned above, one factor that may explain changes in the composition of the roll call agenda is the sorting out of congressional parties, particularly in the South. Defining unsorted districts as those in which the normal presidential vote is less than 50% for the party of the House member in the district, let us consider the percent of unsorted districts as a predictor of whether partisan or bipartisan legislation reaches a roll call vote. Since the baseline probability of a bill reaching a roll call vote is very small, we will restrict the analysis to those bills that were classified as important by Congressional Quarterly Almanac (CQ). The analysis is run at the bill-level where the dependent variable is whether a bill received a roll call vote (as determined by merging data from the Policy Agendas Project with the original data in this paper). The independent variables are bipartisan cosponsorship, defined as 1 for bills where at least 20% of

¹⁵ Full listing of conditional probabilities is available in the online appendix. This analysis includes all cosponsored bills rather than just 'important' bills.

the cosponsors are from the party opposite the party of the sponsor; and the percent of unsorted districts, measured as the percent of districts in each Congress where the normal presidential vote is less than 50%. Table 2 presents the results. Whereas the probability of receiving a roll call vote increases for bipartisan legislation as the percent of unsorted districts increases, the probability decreases for partisan legislation. As the percent of unsorted districts increases from 20% to 50% (the range of the data), the predicted probability of a bill with bipartisan cosponsors receiving a roll call vote increases from 0.37 to 0.51. Across the same range, the predicted probability that a bill with partisan cosponsors receives a roll call vote declines from 0.48 to 0.44.

Conclusions

This research finds that although bipartisan cooperation and electoral responsiveness have declined in patterns of roll call votes, the same is not true for members' use of bill cosponsorship coalitions. Rather, bipartisanship persists in bill cosponsorship coalitions, the stage where legislators are the most independent. While bipartisan bills may not outnumber partisan bills, bipartisanship in cosponsorship coalitions is nearly as common today as it was in the early 1970s. This divergence between roll call voting and bill cosponsorship coalitions can be explained, in part, by changes in the selection of bills to face roll call votes. The findings presented in this paper suggest that sorting and overlap of party members is one possible explanation for changes in agenda strategy. Subsequent work should examine multiple causes, including divided government and the size of majority seat shares, as well as why bipartisan legislation faced a brief resurgence in voting in the late 1990s.

Although partisan behavior has always been a feature of Congress, the main finding of this work is that the relative frequency of partisan relative to bipartisan behavior has not

dramatically changed in bill cosponsorship. While the decline of bipartisanship, and of weakened representation, in roll call voting is certainly important both for its normative implications and for the understanding of legislative behavior, including bill cosponsorship coalitions cautions us against an all-or-nothing account of bipartisan cooperation. When considering bill cosponsorship patterns, Congress has remained representative of, and relatively responsive to, the more moderate public. Subsequent work should examine more broadly how members transmit information about their legislative behavior to constituents, and whether bipartisan cosponsorship can make up for a partisan voting record.

These findings add to the growing literature that cautions against using aggregate roll call data as the basis of all legislative behavior measures (e.g. Roberts 2007; Clinton 2007; Carrubba et al. 2008). With respect to academic work on preferences, polarization, and party power, this research points out the importance of looking beyond roll call votes to explore these relationships. Beyond the discussion of legislative organization and legislative behavior, my findings add to the growing literature on polarization and party strength. While scholars have noted the differences between elite and mass polarization, and have focused on similarities between citizens of red states and blue states (Fiorina et al. 2004), my findings suggest this may occur at the congressional level as well when cosponsorship is considered. The appearance of growing partisan cohesion and polarization reflects how political parties utilize the congressional agenda. Like Cox and McCubbins (2005), I find that the party's ability to select issues for roll call votes has important consequences. In this case, the probability that legislation with bipartisan support faces a roll call vote has varied over time. In effect, the political parties are contributing to the level of party polarization that is observed by the public by selecting which bills face roll call votes.

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Table 1: Correlation of Legislative Behavior with the Normal Presidential Vote

Congress	Bipartisan Cosponsorship	Party Unity Support Score
93	0.03	0.43
94	0.13	0.41
95	0.04	0.43
96	-0.12	0.38
97	-0.22	0.27
98	-0.02	0.13
99	0.03	0.13
100	0.07	0.11
101	-0.16	0.16
102	-0.16	0.29
103	-0.55	0.45
104	-0.46	0.46
105	-0.55	0.36
106	-0.65	0.39
107	-0.63	0.36
108	-0.56	0.38

Source: Calculated by author from cosponsorship matrices provided by James Fowler and roll call data from Rohde (2004). Normal presidential vote was calculated by the author, using data from David Brady.

Table 2: Probability of a Roll Call Vote (CQ Significant Legislation, 1973-2004)

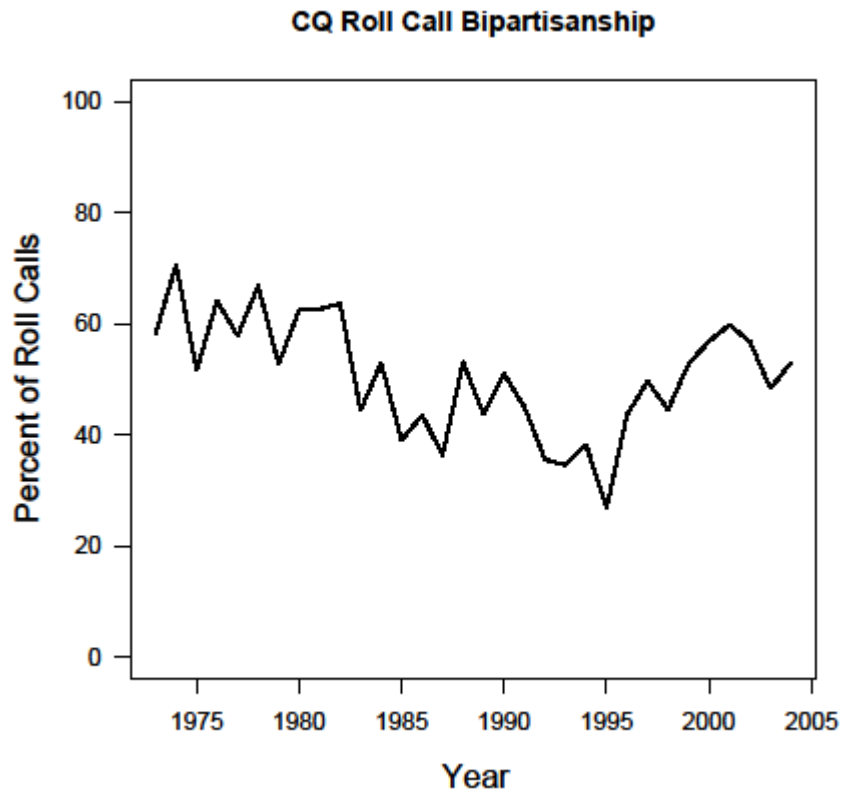
	Model 1
<i>Intercept</i>	-0.00452 (0.229)
<i>Bipartisan Cosponsorship</i>	-0.85** (0.291)
<i>Percent Unsorted Districts</i>	-0.00449 (0.00631)
<i>Bipartisan x Percent Unsorted</i>	0.0225** (0.00791)
<i>N</i>	3313
<i>Log Likelihood</i>	-2274

Standard errors in parentheses.

[^] $p < 0.1$, * $p < 0.05$, ** $p < 0.01$, *** $p < 0.001$.

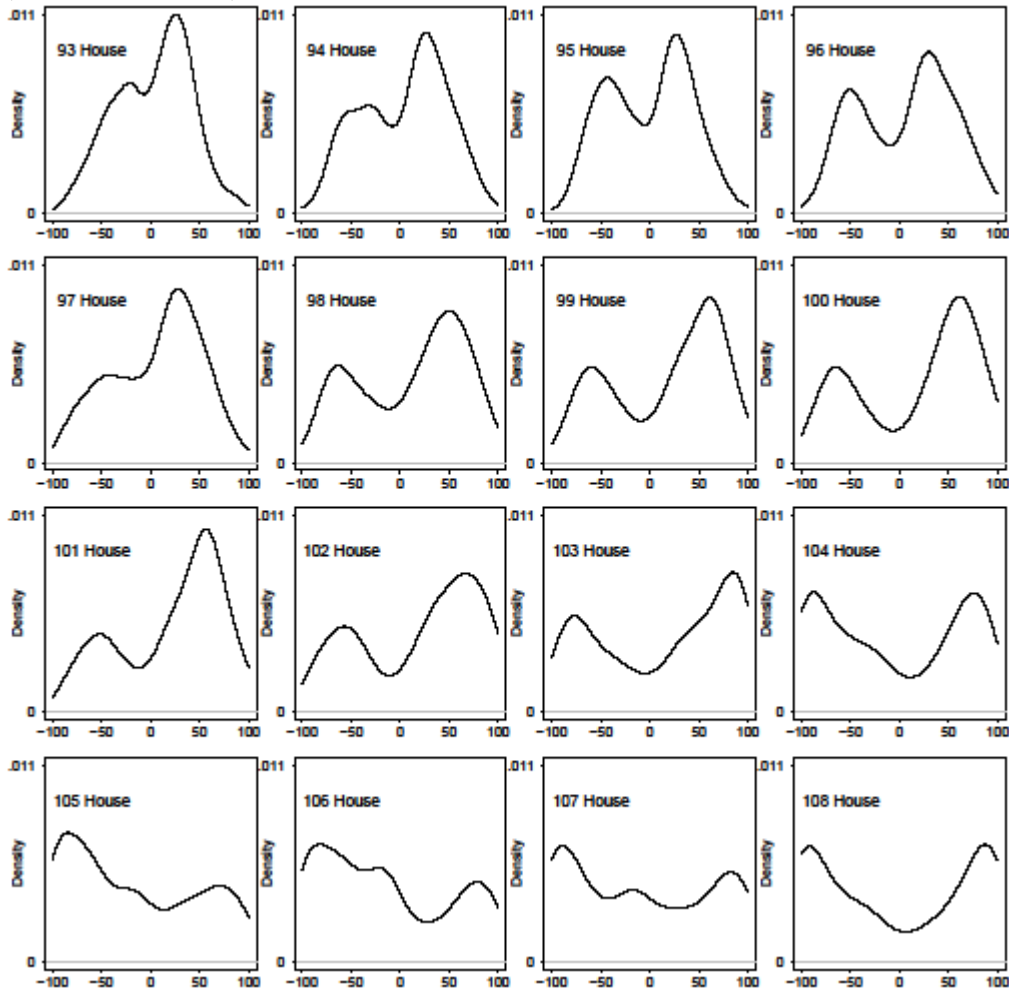
Source: Calculated by author from cosponsorship matrices provided by James Fowler and roll call data provided by Rohde (2004) and the Policy Agenda Project (data were originally collected by Frank R. Baumgartner and Bryan D. Jones, with the support of National Science Foundation grant number SBR 9320922, and were distributed through the Department of Political Science at the University of Washington. Neither NSF nor the original collectors bear responsibility for the analysis reported here.).

Figure 1: CQ Roll Call Bipartisanship Measures (House of Representatives, 1973-2004)



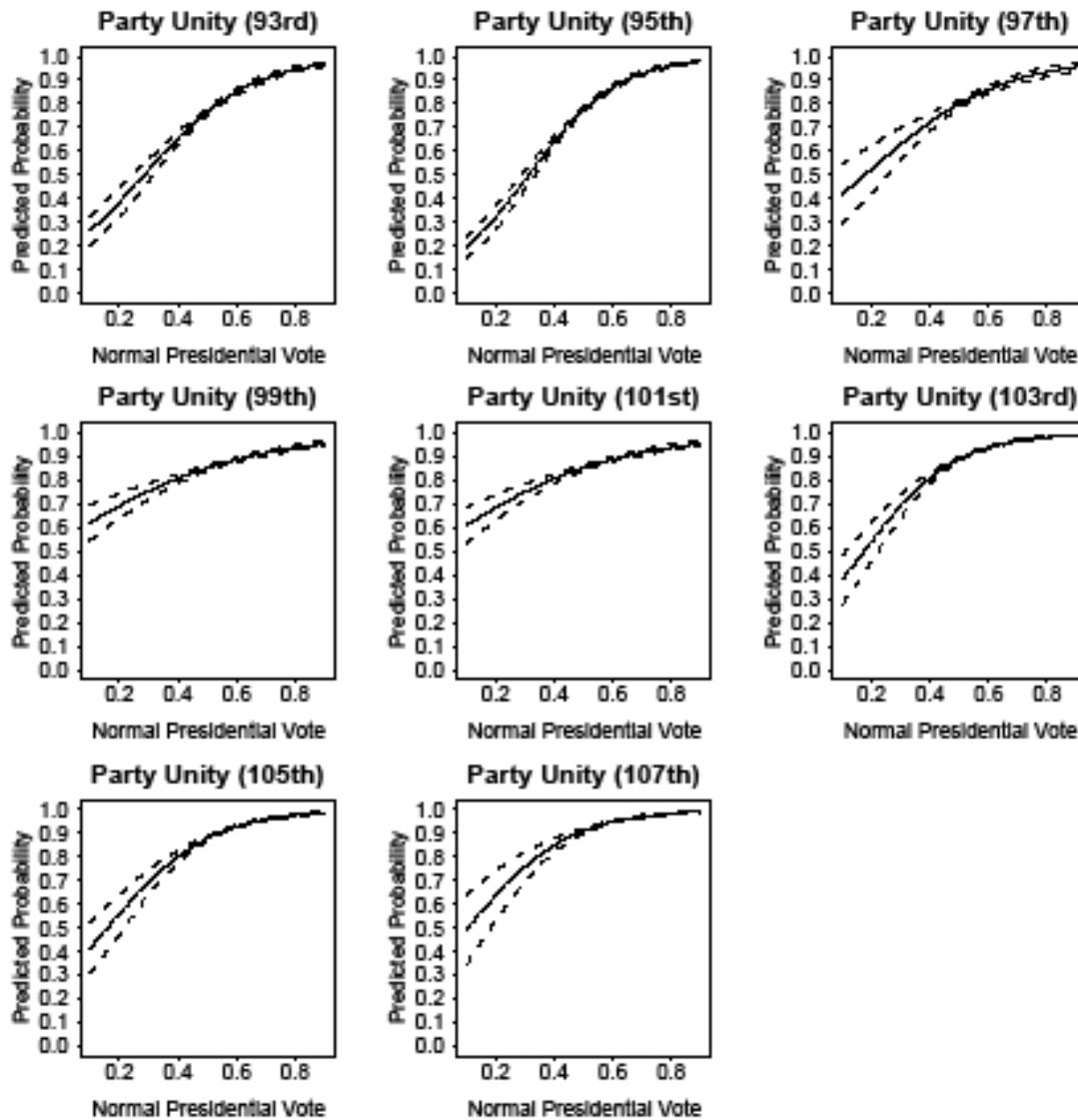
Source: Congressional Quarterly Almanac. Regressing the percent of bipartisan roll call votes on time produces a coefficient of -0.45 ($p < 0.05$). When only final passage roll call votes are included this increases to -0.85 ($p < 0.001$).

**Figure 2: Difference in Percent of Party Voting Yea (% of Democrats - % of Republicans)
(All House Votes)**



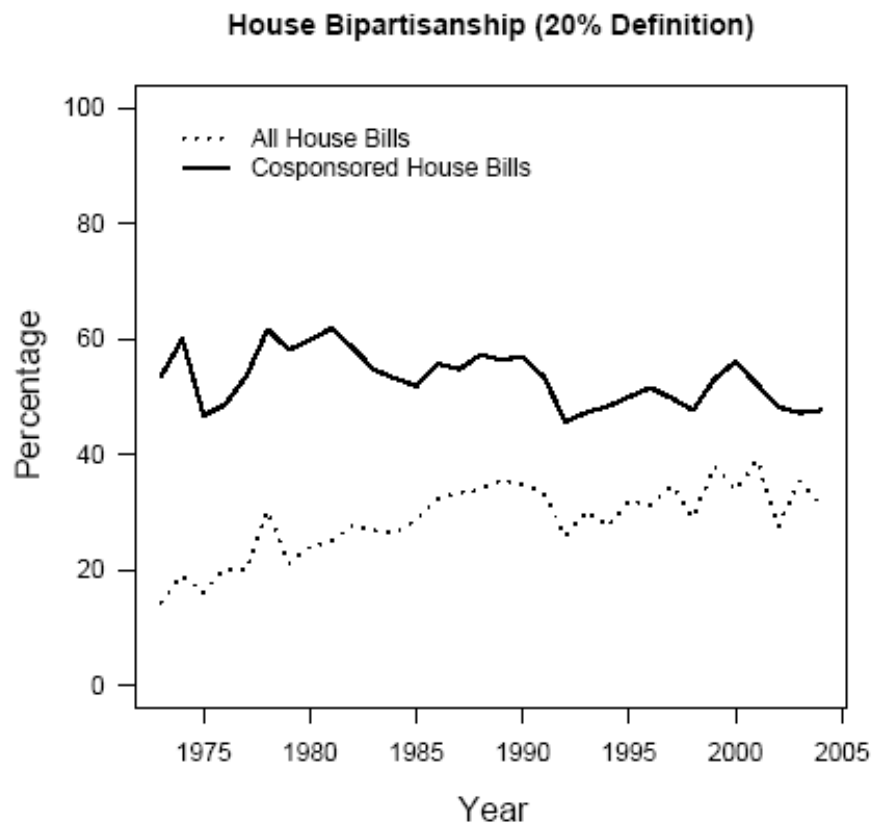
Source: Calculated by the author from Rohde (2004) data. Regressing the mean of the absolute value of the party difference score on time produces a coefficient of 1.16 ($p < 0.01$). When only final passage roll call votes are included this coefficient is 1.11 ($p < 0.001$).

Figure 3: Predicted Probabilities of Partisanship from Quasi-Binomial (Party Unity Support Score, Selected Congresses)



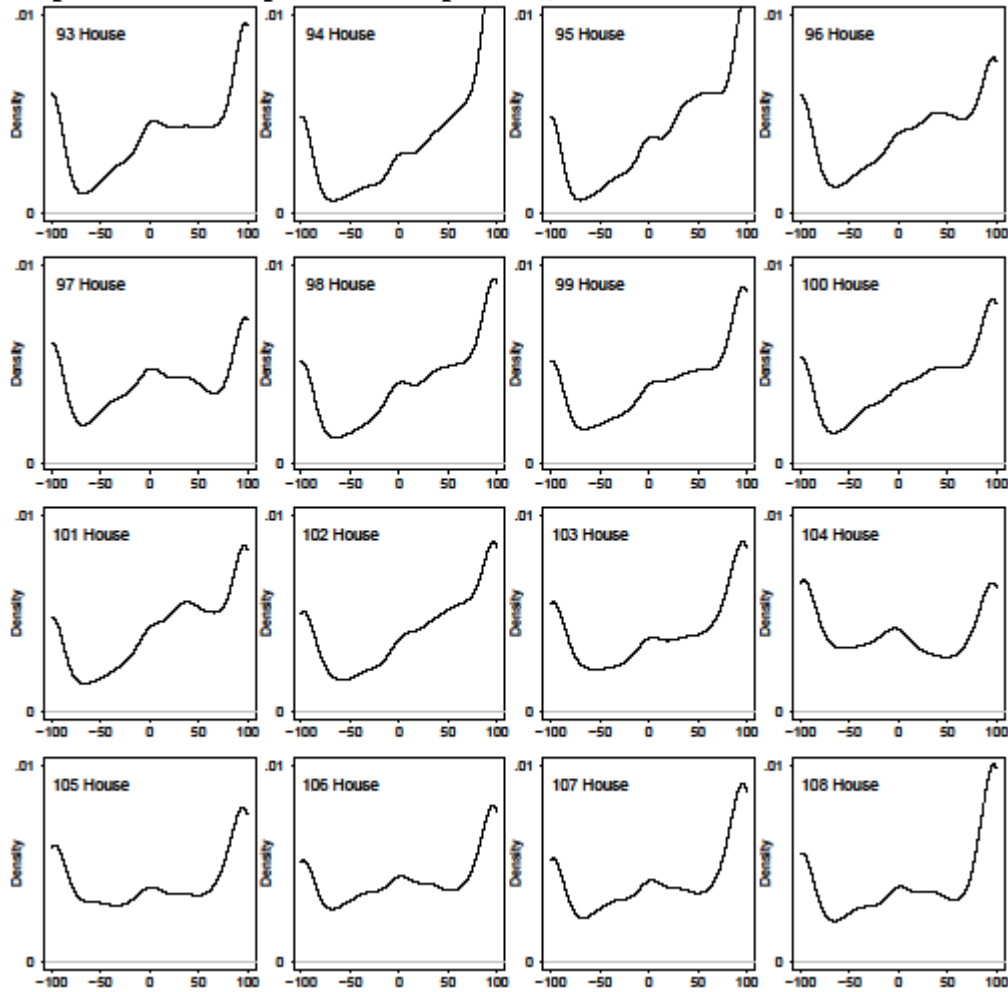
Source: Calculated by author. Full regression model available in online appendix. Normal presidential vote was calculated by the author, using data from David Brady.

Figure 4: Bipartisan Cosponsorship (At least 20% of Cosponsors from Party Opposite the Party of Sponsor, House of Representatives, 1973-2004)



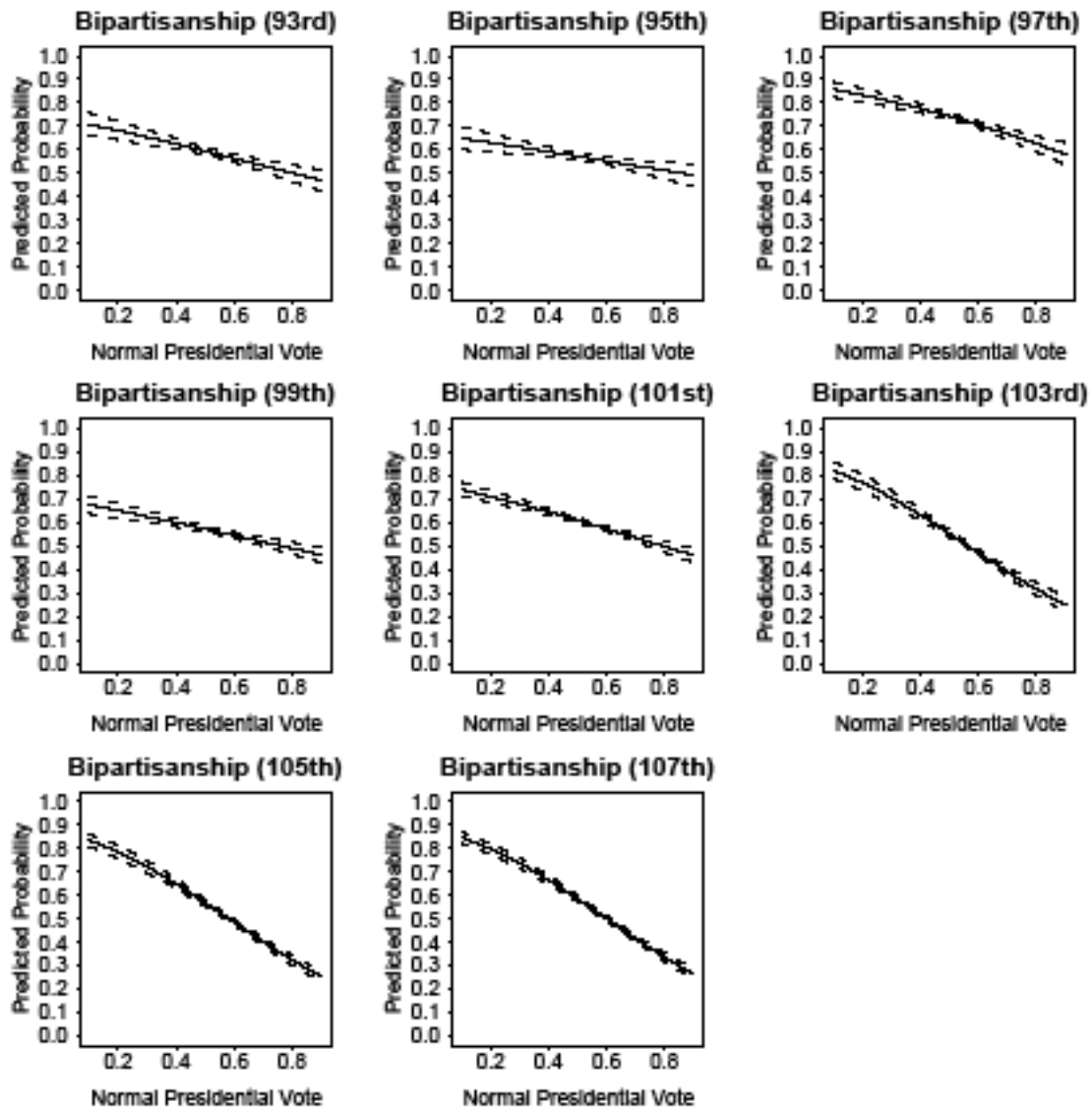
Source: Calculated by author from cosponsorship matrices provided by James Fowler. Regressing the percent of cosponsored bills that are bipartisan on time produces a coefficient of -0.25 ($p < 0.01$). As a comparison, if the measure of bipartisanship is the mean of the percent of cosponsors opposite the party of the bill sponsor, the coefficient on time is -0.17 ($p < 0.001$).

Figure 5: Difference in Party Support in Cosponsorship Coalitions (% Democratic Cosponsors - % Republican Cosponsors)



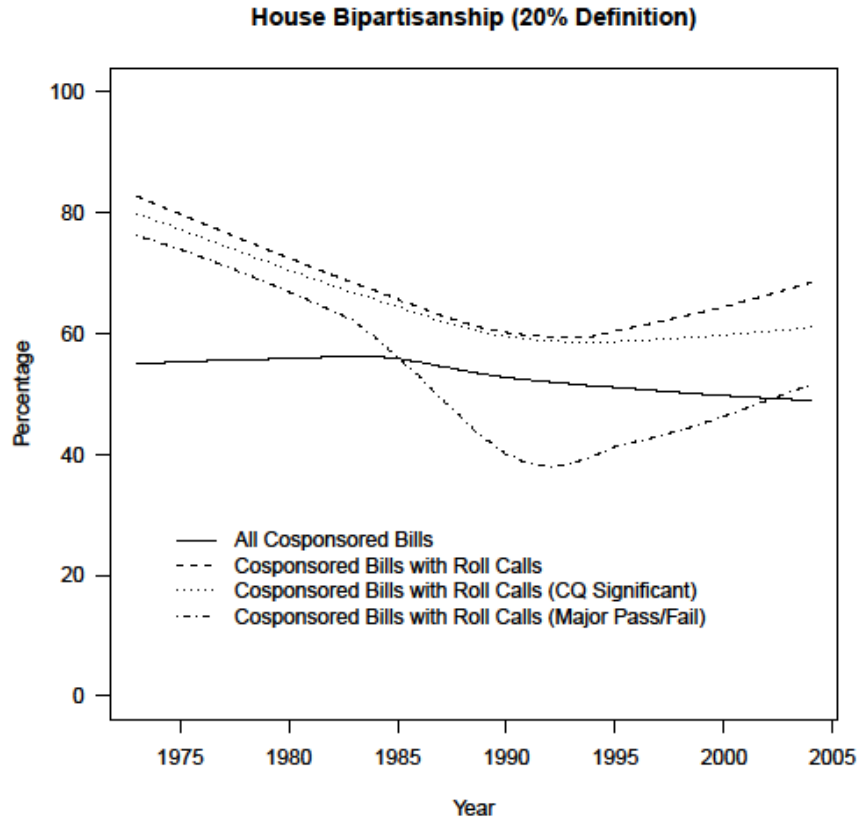
Source: Calculated by author from cosponsorship matrices provided by James Fowler. Regressing the mean of the absolute value of the party scale on time produces a coefficient of -0.02 ($p = 0.72$).

Figure 6: Predicted Probabilities of Bipartisanship from Quasi-Binomial (Bipartisan Cosponsorship, Selected Congresses)



Source: Calculated by author. Full regression model available in online appendix.

Figure 7: Bipartisan Cosponsorship by Stage of Policy Making (Percentage of Cosponsored Bills That Are Bipartisan by 20% Definition, Smoothed)



Source: Calculated by author from cosponsorship matrices provided by James Fowler and roll call data provided by Rohde (2004) and the Policy Agenda Project (data were originally collected by Frank R. Baumgartner and Bryan D. Jones, with the support of National Science Foundation grant number SBR 9320922, and were distributed through the Department of Political Science at the University of Washington. Neither NSF nor the original collectors bear responsibility for the analysis reported here.).

Online Appendix

Appendix Table 1: Full Model Specification of Responsiveness by Members of Congress

	Cosponsorship	Party Unity
<i>Intercept</i>	1.44 ^{***} (0.144)	-2.05 ^{***} (0.209)
<i>Majority Party Member</i>	-0.476 ^{***} (0.013)	0.764 ^{***} (0.0249)
<i>Normal Presidential Vote in District</i>	-0.0127 ^{***} (0.0025)	0.0551 ^{***} (0.00379)
<i>Female</i>	-0.217 ^{***} (0.0186)	-0.0146 (0.0411)
<i>Age</i>	0.00127 [*] (0.000634)	-0.00421 ^{***} (0.00108)
<i>Number of Congresses Served</i>	-0.00338 [^] (0.00192)	-0.0149 ^{***} (0.00318)
<i>House Leadership</i>	-0.386 ^{***} (0.0725)	0.9 ^{***} (0.143)
<i>94th Congress</i>	-1.1 ^{***} (0.184)	0.056 (0.252)
<i>95th Congress</i>	-0.32 [^] (0.191)	-0.513 [^] (0.264)
<i>96th Congress</i>	0.329 [^] (0.196)	-0.341 (0.275)
<i>97th Congress</i>	0.954 ^{***} (0.212)	0.81 [*] (0.375)
<i>98th Congress</i>	-0.343 [^] (0.185)	1.59 ^{***} (0.312)
<i>99th Congress</i>	-0.172 (0.174)	1.75 ^{***} (0.274)
<i>100th Congress</i>	-0.273 (0.17)	1.97 ^{***} (0.278)
<i>101st Congress</i>	0.188 (0.168)	1.69 ^{***} (0.277)
<i>102nd Congress</i>	0.0661 (0.168)	1.38 ^{***} (0.271)
<i>103rd Congress</i>	0.839 ^{***} (0.201)	0.44 (0.331)
<i>104th Congress</i>	0.401 [^] (0.215)	-0.00238 (0.318)
<i>105th Congress</i>	0.938 ^{***} (0.184)	0.645 [^] (0.329)
<i>106th Congress</i>	1.35 ^{***} (0.177)	0.13 (0.357)
<i>107th Congress</i>	1.03 ^{***} (0.175)	0.951 [*] (0.404)
<i>108th Congress</i>	0.539 ^{**}	1.58 ^{***}

	(0.172)	(0.348)
<i>109th Congress</i>	0.248	1.27 ^{***}
	(0.178)	(0.346)
<i>94th Congress x Normal Vote</i>	0.00855 [*]	0.00238
	(0.00335)	(0.00489)
<i>95th Congress x Normal Vote</i>	0.00457	0.0111 [*]
	(0.00347)	(0.00509)
<i>96th Congress x Normal Vote</i>	-0.00247	0.0112 [*]
	(0.00354)	(0.00531)
<i>97th Congress x Normal Vote</i>	-0.00556	-0.0122 [^]
	(0.00374)	(0.00697)
<i>98th Congress x Normal Vote</i>	0.00462	-0.024 ^{***}
	(0.00329)	(0.00582)
<i>99th Congress x Normal Vote</i>	0.00163	-0.0238 ^{***}
	(0.00309)	(0.0051)
<i>100th Congress x Normal Vote</i>	0.00462	-0.0253 ^{***}
	(0.00303)	(0.00518)
<i>101st Congress x Normal Vote</i>	-0.00249	-0.0225 ^{***}
	(0.00298)	(0.00517)
<i>102nd Congress x Normal Vote</i>	-0.00433	-0.0135 ^{**}
	(0.00298)	(0.00512)
<i>103rd Congress x Normal Vote</i>	-0.0202 ^{***}	0.00928
	(0.00354)	(0.00625)
<i>104th Congress x Normal Vote</i>	-0.0117 ^{**}	0.0167 ^{**}
	(0.00377)	(0.00596)
<i>105th Congress x Normal Vote</i>	-0.0209 ^{***}	0.00265
	(0.00317)	(0.00604)
<i>106th Congress x Normal Vote</i>	-0.0252 ^{***}	0.0127 [^]
	(0.00307)	(0.00656)
<i>107th Congress x Normal Vote</i>	-0.0214 ^{***}	0.00324
	(0.003)	(0.00738)
<i>108th Congress x Normal Vote</i>	-0.0159 ^{***}	-0.00358
	(0.00296)	(0.00644)
<i>109th Congress x Normal Vote</i>	-0.013 ^{***}	-0.000174
	(0.00306)	(0.00634)
<hr/>		
<i>N</i>	7385	7412
<i>Log Likelihood</i>	—	—

Standard errors in parentheses.

[^] $p < 0.1$, * $p < 0.05$, ** $p < 0.01$, *** $p < 0.001$.

Appendix Table 2: Cosponsorship Summary Statistics

Congress	Number of House Bills	Proportion of Bills Cosponsored	Median # Cosponsors (all bills)	Median # Cosponsors (> 0 cosponsors)	Proportion of bills with 0-10 Cosponsors	Proportion of bills with >100 Cosponsors
93	17,690	0.28	0.00	6.00	0.60	0.00
94	15,863	0.36	0.00	6.00	0.61	0.00
95	14,414	0.40	0.00	6.00	0.59	0.00
96	8,455	0.37	0.00	6.00	0.57	0.02
97	7,457	0.42	0.00	7.00	0.59	0.03
98	6,442	0.49	0.00	7.00	0.56	0.04
99	5,753	0.56	1.00	9.00	0.51	0.04
100	5,585	0.60	1.00	9.00	0.47	0.05
101	5,977	0.62	2.00	9.00	0.48	0.05
102	6,212	0.60	2.00	10.00	0.49	0.05
103	5,310	0.62	2.00	10.00	0.50	0.04
104	4,344	0.63	2.00	10.00	0.51	0.05
105	4,874	0.67	3.00	11.00	0.47	0.05
106	5,681	0.67	3.00	11.00	0.48	0.06
107	5,767	0.68	3.00	11.00	0.47	0.05
108	5,431	0.72	4.00	11.00	0.49	0.06

Source: Calculated by author from cosponsorship matrices provided by James Fowler.

Appendix Table 3: Regression of Bipartisanship on Time (OLS, 1973-2004)

	Mean Party Score (Cosponsorship)	20% Definition (Cosponsorship)	% Bipartisan (Roll Calls)	% Bipartisan (Final Passage Roll Calls)	Mean Party Score (Roll Call)
<i>Intercept</i>	65.9*** (1.06)	57.2*** (1.49)	58*** (3.53)	73.9*** (2.7)	33.8*** (1.67)
<i>Time Trend</i>	-0.0201 (0.0562)	-0.248** (0.0788)	-0.449* (0.187)	-0.845*** (0.143)	1.16*** (0.0885)
<i>N</i>	32	32	32	32	32
<i>R</i> ²	0.00426	0.248	0.162	0.538	0.851
<i>Adjusted R</i> ²	-0.0289	0.223	0.134	0.523	0.846

Standard errors in parentheses.

[^] $p < 0.1$, * $p < 0.05$, ** $p < 0.01$, *** $p < 0.001$.

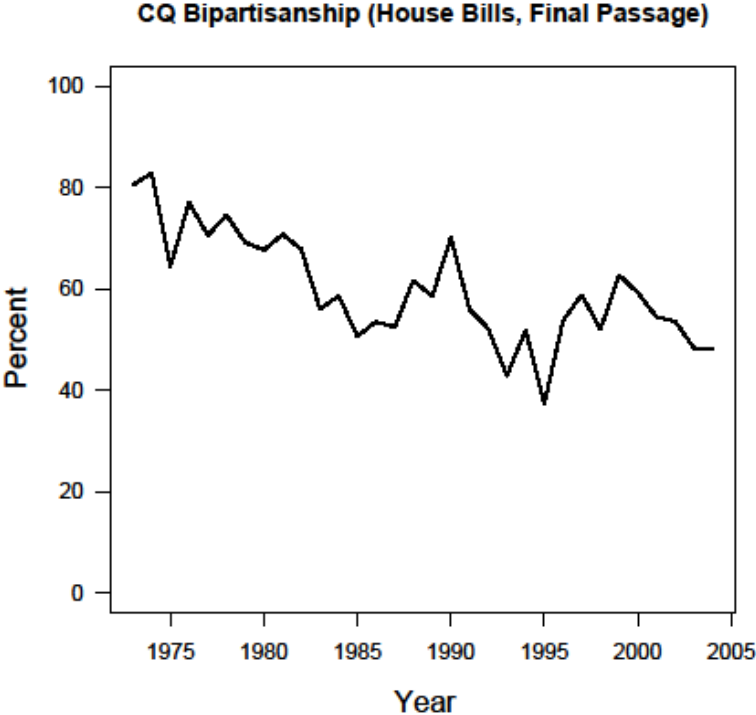
Source: Calculated by author from cosponsorship matrices provided by James Fowler and roll call data provided by Rohde (2004) and the Policy Agenda Project (data were originally collected by Frank R. Baumgartner and Bryan D. Jones, with the support of National Science Foundation grant number SBR 9320922, and were distributed through the Department of Political Science at the University of Washington. Neither NSF nor the original collectors bear responsibility for the analysis reported here.).

Appendix Table 4: Conditional Probability of Reaching a Roll Call Given Being a Partisan/Bipartisan (20% Definition) Bill

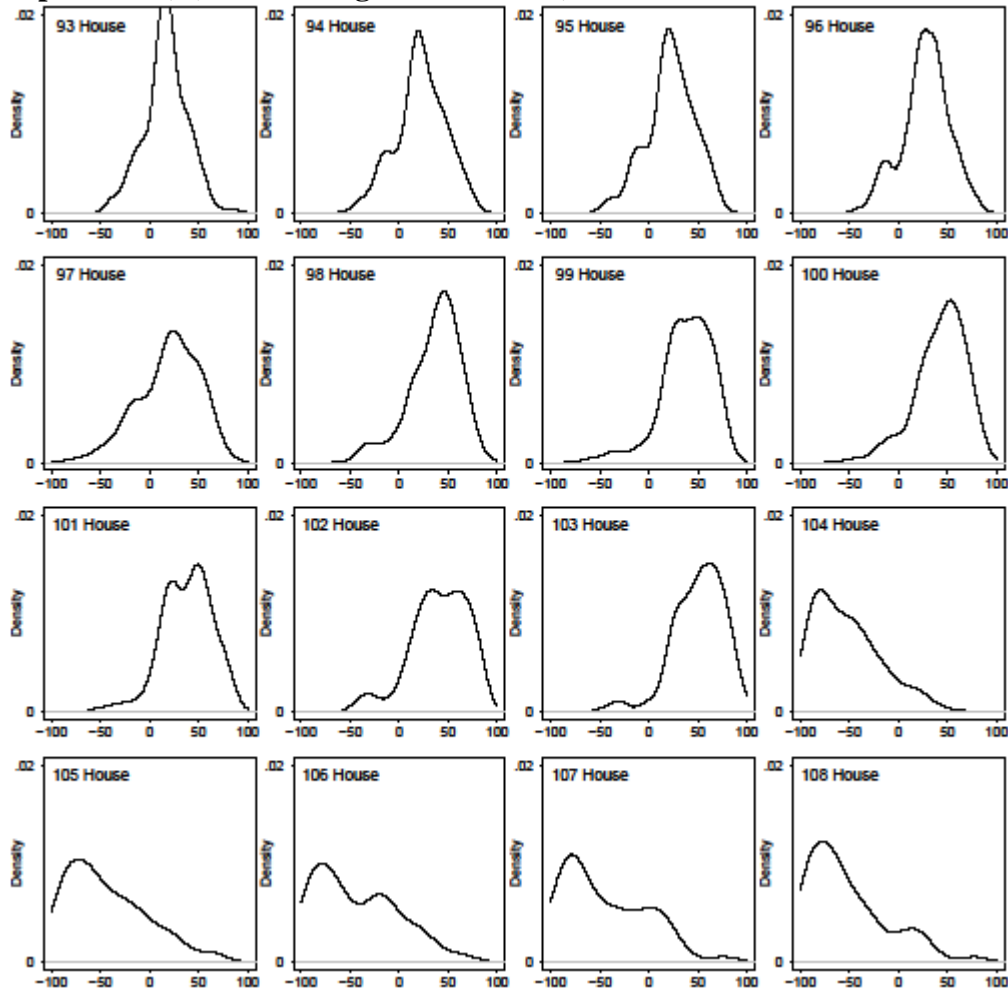
Congress	Roll Call Partisan	Roll Call Bipartisan
93	0.01	0.06
94	0.02	0.07
95	0.02	0.06
96	0.04	0.07
97	0.02	0.05
98	0.05	0.06
99	0.03	0.04
100	0.03	0.05
101	0.03	0.03
102	0.03	0.03
103	0.03	0.03
104	0.04	0.05
105	0.04	0.06
106	0.03	0.08
107	0.03	0.07
108	0.03	0.07

Source: Calculated by author from cosponsorship matrices provided by James Fowler and roll call data provided by Rohde (2004) and the Policy Agenda Project (data were originally collected by Frank R. Baumgartner and Bryan D. Jones, with the support of National Science Foundation grant number SBR 9320922, and were distributed through the Department of Political Science at the University of Washington. Neither NSF nor the original collectors bear responsibility for the analysis reported here.).

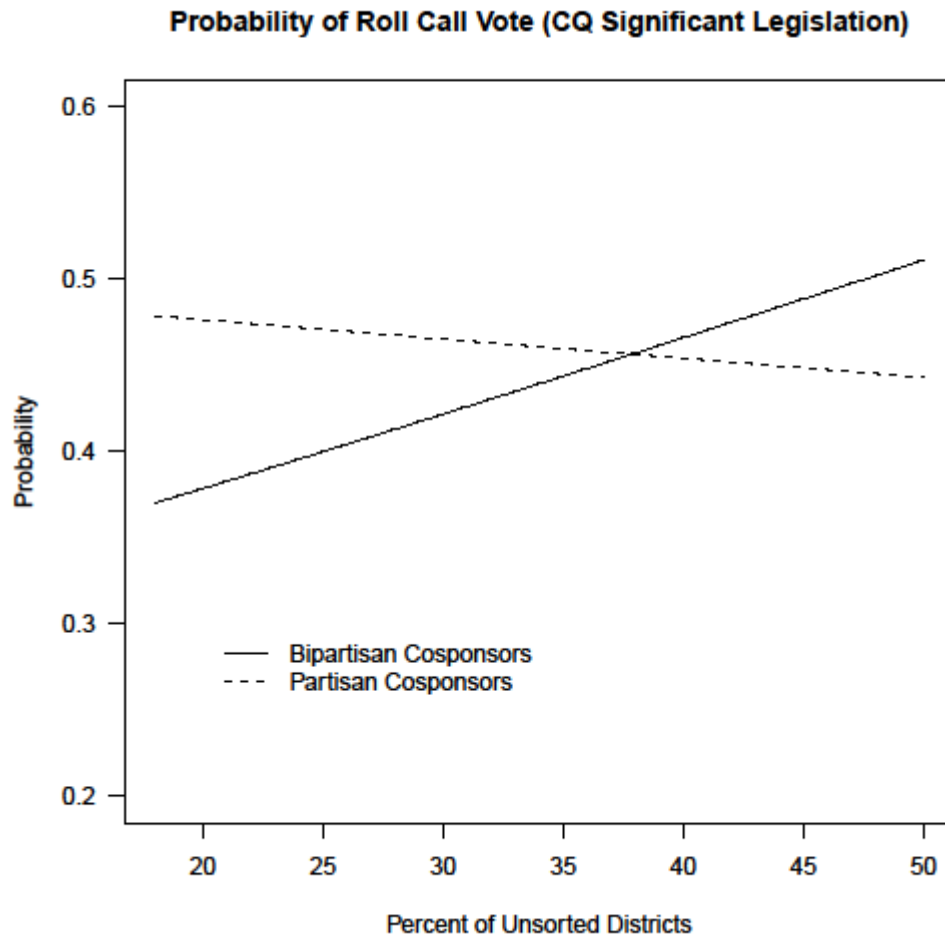
Appendix Figure 1: CQ Roll Call Bipartisanship Measures (House of Representatives, Final Passage Votes, 1973-2004)



Appendix Figure 2: Difference in Percent of Party Voting Yea (% of Democrats - % of Republicans) (Final Passage House Votes)



Appendix Figure 3: Predicted Probabilities from Table 4



Source: Calculated by author from cosponsorship matrices provided by James Fowler and roll call data provided by Rohde (2004) and the Policy Agenda Project (data were originally collected by Frank R. Baumgartner and Bryan D. Jones, with the support of National Science Foundation grant number SBR 9320922, and were distributed through the Department of Political Science at the University of Washington. Neither NSF nor the original collectors bear responsibility for the analysis reported here.).